

Lecture 17: May 6, 2005

Instructor: László Babai

Scribe: Hariharan Narayanan

NOTE: Change in Monday's TA schedule; no change Tuesday and Thursday.

TA SCHEDULE: TA sessions are held in Ryerson-255, Monday 7:30–8:30 ,

Tuesday and Thursday 5:30–6:30pm.

INSTRUCTOR'S EMAIL: laci@cs.uchicago.edu

TA's EMAIL: hari@cs.uchicago.edu, raghav@cs.uchicago.edu

The second moment method

Lemma 17.1 (The second moment method) *If $X \geq 0$ is a random variable not identically zero, then $P(X = 0) \leq \frac{\text{var}(X)}{E(X)^2}$.*

Recall Chebychev's inequality. $P(|X - E(X)| \geq a) \leq \frac{\text{var}(X)}{a^2}$

Let $a := E(X) > 0$.

$$\begin{aligned} P(X = 0) &\leq P(|X - E(X)| \geq E(X)) \\ &\leq \frac{\text{var}(X)}{a^2} \\ &= \frac{\text{var}(X)}{(E(X))^2}. \end{aligned}$$

$$\text{var}(X) = \sum_i \text{var}(X_i) + 2 \sum_{i < j} \text{cov}(Y_i, Y_j) < 2 \sum_{i \leq j} \text{cov}(Y_i, Y_j).$$

$$\text{cov}(S, T) = E(ST) - E(S)E(T).$$

□

Let X be the number of runs of k heads in a sequence of n coin flips.

Theorem 17.2 1. $n = n(k)$. If $\frac{n(k)}{2^k} \rightarrow 0$ then $P(X \neq 0) \rightarrow 0$.

2. If $\frac{n(k)}{2^k} \rightarrow \infty$ then $P(X \neq 0) \rightarrow 1$. (“Almost surely, there exists a run of k heads”).
 $Y = Y_1 + \dots + Y_{n-k+1}$. Y_i indicates the event that the $i^{\text{th}}, \dots, (i+k-1)^{\text{th}}$ coin flips come out heads.

Proof:

$E(Y_i) = P(Y_i = 1) = 1/2^k$. Therefore $E(X) = \sum_{i=1}^{n-k+1} E(Y_i) = \frac{n-k+1}{2^k}$. To prove (1.), observe that $P(X \neq 0) = P(X \geq 1) \leq E(X)$. To prove (2.), we use the *second moment method*.

Claim 17.3 $\text{var}(X) < 5(E(X))$

Proof of Theorem 17.2 using Claim 17.3 :

$\text{var}(X) < cE(X)$.

$P(X = 0) \leq \frac{\text{var}(X)}{E(X)^2} < \frac{c}{E(X)} \rightarrow 0$.

□

Proof of Claim 17.3:

Suppose $|j - i| = k - \ell$, $\ell \geq 1$ being the “size of the overlap”.

$$\begin{aligned} \text{cov}(Y_i Y_j) &= E(Y_i Y_j) - E(Y_i)E(Y_j) \\ &< E(Y_i Y_j) \\ &< 2^{-(2k-\ell)} \end{aligned}$$

Therefore

$$\begin{aligned} \sum_{i \leq j} \text{cov}(Y_i, Y_j) &< \sum_{\ell=0}^k \frac{n-k+1}{2^{2k-\ell}} \\ &= \frac{n-k+1}{2^{2k}} \sum_{\ell=0}^k 2^\ell \\ &< \frac{n-k+1}{2^{2k}} \sum_{\ell=0}^k 2^\ell \\ &< 2 \frac{n-k+1}{2^k} \\ &= 2E(X) \end{aligned}$$

Therefore, $\text{var}(X) \leq 2 \sum_{i \leq j} \text{cov}(Y_i, Y_j) < 4E(X)$. □

Theorem 17.4 For almost all graphs on n vertices, $\alpha(G) \sim 2 \log_2(n)$.

We already know : $\lesssim 2 \log_2 n$. We need: $\gtrsim 2 \log_2 n$.

Method: 2^{nd} -Moment method. (Hint: X = the number of anticliques of size k .)

Definition 17.5 $\tau(\mathcal{H})$ is the covering number of a k -uniform hypergraph with m edges and n vertices.

i. e. $\tau(\mathcal{H}) = \min |T| : T \text{ such that } (\forall i)(T \cap A_i \neq \emptyset)$.

Claim 17.6 $\tau(\mathcal{H}) < 1 + n/k \ln(m)$.

Claim 17.7 *If T is a random subset of size t , $t \geq n/k \ln m$, then with probability greater than 0, T is a cover. $P(A_i \text{ is not hit by } T) = \frac{\binom{n-k}{t}}{\binom{n}{t}}$.*

Proof:

$$P(A_i \text{ is not hit by } T) = \frac{\binom{n-k}{t}}{\binom{n}{t}}.$$

$$P((\exists i) A_i \text{ is not hit by } T) < m \frac{\binom{n-k}{t}}{\binom{n}{t}} \leq 1.$$

□

Exercise 17.8

$$\begin{aligned} \frac{\binom{n-k}{t}}{\binom{n}{t}} &\leq \left(\frac{n-k}{n}\right)^t \\ &= (1 - k/n)^t < e^{-kt/n} \\ &\leq e^{-\ln m} = 1/m \end{aligned}$$